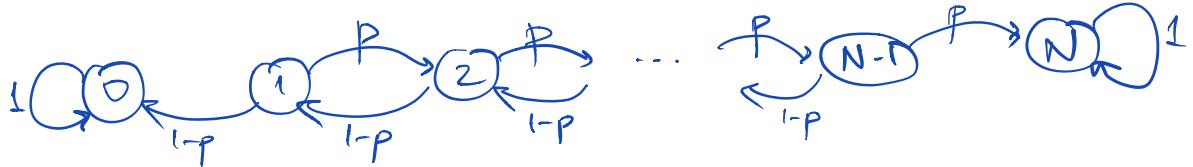


Last time: Gambler's ruin problem



Q: What is  $P(\text{lose} \mid X_0 = n) =: \bar{p}(n)$ .

We found a recurrence relation:

$$\begin{aligned}\bar{p}(n) &= \bar{p}(n+1) \cdot p + \bar{p}(n-1) \cdot (1-p) \\ \bar{p}(0) &= 1 \quad \left. \begin{array}{l} \text{boundary / initial} \\ \bar{p}(N) = 0 \end{array} \right\} \text{conditions}\end{aligned}$$

Characteristic equation:

$$x^2 - \frac{1}{p}x + \frac{(1-p)}{p} = 0$$

$$\text{solutions } y_1 = 1, \quad y_2 = \frac{1-p}{p}$$

Case 1:  $y_1 + y_2$ , i.e.  $p + \frac{1}{2}$ .

$$\text{Then, } \bar{p}(n) = \alpha y_1^n + \beta y_2^n = \alpha + \beta \cdot \left(\frac{1-p}{p}\right)^n$$

Substitute initial conditions:

$$\left\{ \begin{array}{l} \bar{p}(0) = 1 \\ \bar{p}(N) = 0 \end{array} \right. \Leftrightarrow \left\{ \begin{array}{l} \alpha + \beta = 1 \\ \alpha + \beta \left(\frac{1-p}{p}\right)^N = 0 \end{array} \right. \Leftrightarrow \left\{ \begin{array}{l} \beta = 1 - \alpha \\ \alpha + (1-\alpha) \left(\frac{1-p}{p}\right)^N = 0 \end{array} \right.$$

$$\begin{aligned} &\Leftrightarrow \left\{ \begin{array}{l} \alpha = \frac{\left(\frac{1-p}{p}\right)^N}{\left(\frac{1-p}{p}\right)^N - 1} \\ \beta = 1 - \alpha = \frac{-1}{\left(\frac{1-p}{p}\right)^N - 1} \end{array} \right. \\ &\Rightarrow \bar{p}(n) = \frac{\left(\frac{1-p}{p}\right)^N}{\left(\frac{1-p}{p}\right)^N - 1} - \frac{1}{\left(\frac{1-p}{p}\right)^N - 1} \left(\frac{1-p}{p}\right)^n = \boxed{\frac{\left(\frac{1-p}{p}\right)^N - \left(\frac{1-p}{p}\right)^n}{\left(\frac{1-p}{p}\right)^N - 1}} \end{aligned}$$

Case 2:  $y_1 = y_2$ , i.e.,  $p = \frac{1}{2}$ .

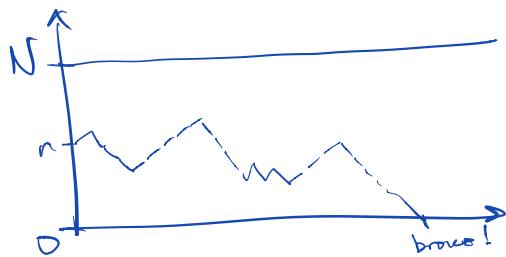
$$\text{Then, } \bar{p}(n) = \alpha y_1^n + \beta n y_1^n = \alpha + \beta n.$$

Substitute initial conditions:

$$\left\{ \begin{array}{l} \bar{p}(0) = 1 \\ \bar{p}(N) = 0 \end{array} \right. \Leftrightarrow \left\{ \begin{array}{l} \alpha = 1 \\ \alpha + N\beta = 0 \end{array} \right. \Leftrightarrow \left\{ \begin{array}{l} \alpha = 1 \\ \beta = -\frac{1}{N} \end{array} \right.$$

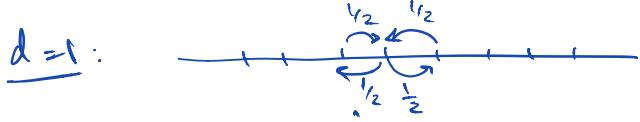
$$\text{Thus, } \bar{p}(n) = \alpha + \beta n = \boxed{1 - \frac{n}{N}}$$

Remark: We can also interpret the gambler's ruin problem as a 1-D random walk with absorbing boundaries s.t. we want to find  $P(\text{hit } 0 \text{ before we hit } N)$



### B. Random walk on $\mathbb{Z}^d$

We consider the uniform r.w. in  $\mathbb{Z}^d$



Question: Is the process recurrent or transient?

We study  $\sum_{n=0}^{\infty} p_{00}^{(n)}$ .

$d=1$ : Starting from 0, we can only come back in an even # of steps.

$p_{00}^{(2n)}$  = Prob to move n times backward and n times forward

$$= \underbrace{\left(\frac{1}{2}\right) \times \left(\frac{1}{2}\right) \times \dots \times \left(\frac{1}{2}\right)}_{2n} \times \underbrace{\binom{2n}{n}}_{\substack{\text{Total # of trajectories} \\ \text{choosing } n \text{ times when the trajectory goes forward.}}} = \frac{\binom{2n}{n}}{2^{2n}}$$

$$\Rightarrow \sum_{n=0}^{+\infty} p_{00}^{(n)} = \sum_{n=0}^{+\infty} \frac{\binom{2n}{n}}{2^{2n}} = \sum_{n=0}^{\infty} \frac{(2n)!}{(n!)^2 2^{2n}}$$

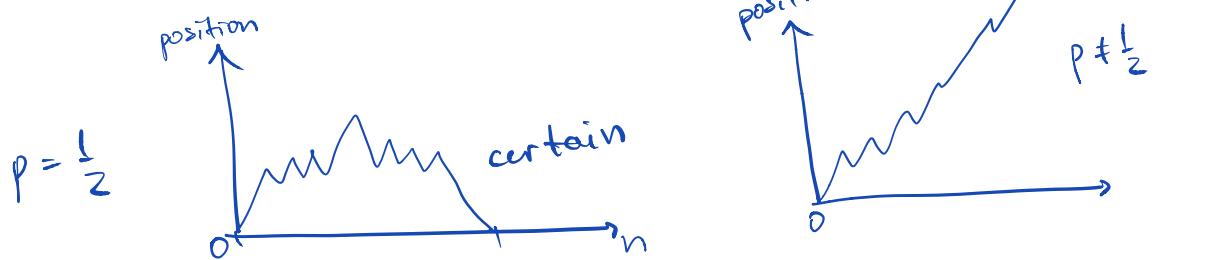
Stirling formula:  $n! \sim \left(\frac{n}{e}\right)^n \sqrt{2\pi n}$

$$\begin{aligned} \Rightarrow \frac{(2n)!}{(n!)^2 2^{2n}} &\approx \frac{(2n)^{2n}}{e^{2n}} \sqrt{2\pi 2n} \left(\frac{n}{e}\right)^{-2n} \frac{1}{2^{2n} n} \cdot \frac{1}{2^{2n}} \\ &\approx \frac{1}{\sqrt{\pi n}} = \left(\frac{1}{\pi n}\right)^{1/2} \end{aligned}$$

$$\Rightarrow \sum_{n=0}^{+\infty} P_{00}^{(n)} = +\infty$$

$\Rightarrow$  The R.W. is recurrent when  $p = \frac{1}{2}$ .

Remark: When  $d=1$  and  $p \neq \frac{1}{2}$ , we can show that R.W. is transient! (asymmetric R.W.)



- For  $d \geq 2$ , we can use a similar approach, but the problem becomes more and more complex as  $d$  increases, so we are going to use a different approach: characteristic functions.

Notation: unit vectors in  $\mathbb{Z}^d$ :  $\vec{e}_j = (0, \dots, 0, \overset{j\text{th position}}{1}, 0, \dots, 0)$   
(e.g.,  $d=3$ :  $\vec{e}_1 = (1, 0, 0)$ ,  $\vec{e}_2 = (0, 1, 0)$ ,  $\vec{e}_3 = (0, 0, 1)$ ).

So, the walk takes  $i^{\text{th}}$  step  $\vec{X}_i$  with probabilities:

$$P(\vec{X}_i = \vec{e}_j) = P(X_i = -\vec{e}_j) = \frac{1}{2d}.$$

let  $\vec{S}_0 = 0$  initial position

$$\begin{aligned}\vec{S}_n &= \text{position after } n \text{ steps} \\ &= \vec{X}_1 + \vec{X}_2 + \dots + \vec{X}_n \quad (n \geq 1).\end{aligned}$$

Then,  $\vec{S}_{n+1} = \vec{S}_n + \vec{X}_{n+1}$ , so

$$P(\vec{S}_{n+1} = \vec{x} + \vec{e}_j \mid \vec{S}_n = \vec{x}) = P(\vec{X}_{n+1} = \vec{e}_j) = \frac{1}{2d},$$

and

$$P(\vec{S}_{n+1} = \vec{x} - \vec{e}_j \mid \vec{S}_n = \vec{x}) = P(\vec{X}_{n+1} = -\vec{e}_j) = \frac{1}{2d}.$$

This defines the transition probabilities for the symmetric R.W. on  $\mathbb{Z}^d$ .

For  $\vec{k} \in \mathbb{R}^d$ ,  $k = (k_1, \dots, k_d)$  we can define  $\Phi_n(\vec{k})$  to be the characteristic function of  $\vec{S}_n$  evaluated at  $\vec{k}$ :

$$\Phi_n(\vec{k}) = \mathbb{E}[e^{i\vec{k} \cdot \vec{S}_n}],$$

where  $\vec{k} \cdot \vec{S}_n$  is the standard dot product in  $\mathbb{R}^d$ , i.e.,

$$\vec{x} \cdot \vec{y} = x_1 y_1 + \dots + x_n y_n.$$

Recall Euler's formula:  
 $e^{i\theta} = \cos\theta + i\sin\theta$ .