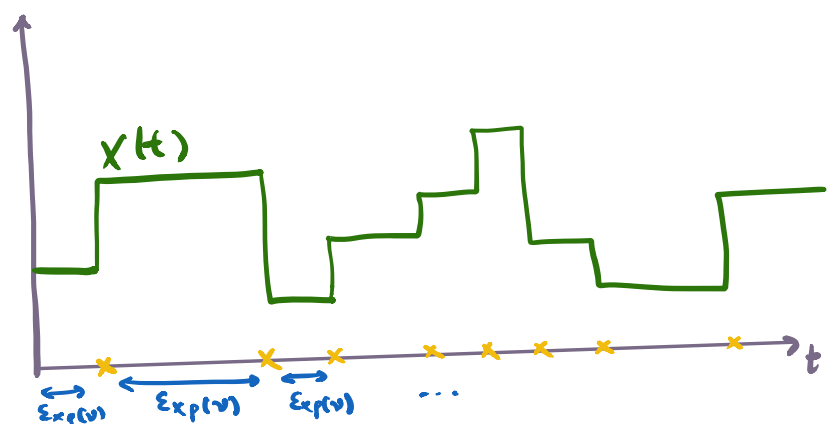


## Uniformization (Section 6.7 in Ross)

- Let  $\{X(t), t \geq 0\}$  be a CTMC for which  $\nu_i = \nu$  for all states  $i$ . (i.e.,  $T_i \sim \text{Exp}(\nu) \forall i$ )

Goal: give a formula for  $P_{ij}(t)$ .

Let  $N(t) = \#$  state transitions by time  $t$ .



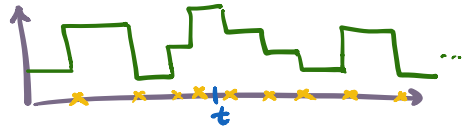
$N(t) = \#$  transitions by time  $t$ .

Since time between 2 state transitions of  $X(t)$  is  $\text{Exp}(\nu)$ , then  $N(t)$  is a Poisson process of rate  $\nu$ !

Transition probabilities  $P_{ij}(t)$  for  $X(t)$  satisfy:

$$\begin{aligned} P_{ij}(t) &= P(X(t) = j | X(0) = i) \\ \text{split according to value of } N(t) &= \sum_{n=0}^{\infty} P(X(t) = j, N(t) = n | X(0) = i) \\ \text{condition on } N(t) &= \sum_{n=0}^{\infty} P(X(t) = j | X(0) = i, N(t) = n) \underbrace{P(N(t) = n | X(0) = i)}_{\sim \text{Poisson}(t\nu)} \\ &= \sum_{n=0}^{\infty} P(X(t) = j | X(0) = i, N(t) = n) e^{-t\nu} \frac{(t\nu)^n}{n!} \end{aligned}$$

What is  $P(X(t)=j | X(0)=i, N(t)=n)$ ?



Since the distribution of time spent in each state  $k$  is the same for all  $k$ , given there were  $n$  transitions:

$$P(X(t)=j | X(0)=i, N(t)=n) = P_{ij}^n,$$

where  $P^n$  is the  $n$ -step transition probability matrix of the embedded chain.

Therefore,

$$P_{ij}(t) = \sum_{n=0}^{\infty} P_{ij}^n e^{-\nu t} \frac{(\nu t)^n}{n!}.$$

This is often useful for approximating  $P_{ij}(t)$  by taking a partial sum.

- How can we turn any CTMC into one with  $\nu_i = \nu$ ?

Consider a CTMC  $X(t)$  for which the  $\nu_i$  are bounded, i.e.,  $\exists \nu$  s.t.  $\nu_i \leq \nu \quad \forall i$ .

When in state  $i$   $X(t)$  leaves at rate  $\nu_i$ . This is equivalent to supposing the transitions occur at rate  $\nu$  (i.e. more often), but

- only the fraction  $\frac{\nu_i}{\nu}$  of these transitions are real (i.e. they leave state  $i$ )
- the remaining  $1 - \frac{\nu_i}{\nu}$  "transition" to state  $i$  itself

Thus,  $X(t)$

- spends  $\text{Exp}(\nu)$  in state  $i$ , and
- transitions to state  $j$  with probability:

$$P_{ij}^* = \begin{cases} 1 - \frac{\nu_i}{\nu} & \text{if } i=j \\ \frac{\nu_i}{\nu} \cdot P_{ij} & \text{if } i \neq j \end{cases}$$

Thus, the transition probabilities can be computed as

$$P_{ij}(t) = \sum_{n=0}^{\infty} P_{ij}^{*n} e^{-\nu t} \frac{(\nu t)^n}{n!},$$

where  $P^{*n}$  is the  $n$ -step transition matrix of a discrete-time MC with transition matrix  $P^*$ .

This technique is called **uniformization**.

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Example (6.11 textbook): Recall the 2-state CTMC

state 0: working, state 1: broken



time to breakdown  $\sim \text{Exp}(\lambda)$

$$\nu_0 = \lambda$$

$$P_{01} = 1$$

time to repair  $\sim \text{Exp}(\mu)$

$$\nu_1 = \mu$$

$$P_{10} = 1$$

Q: Find  $P_{ij}(t)$ .

Solution: Let  $\nu = \lambda + \mu \Rightarrow \nu \geq \lambda, \mu$ .

Uniformized transition probabilities:

$$P_{00}^* = 1 - \frac{\nu_0}{\nu} = \frac{\mu}{\lambda + \mu}; \quad P_{01}^* = \frac{\lambda}{\lambda + \mu}$$

$$p_{10}^* = \frac{\nu_1}{\nu} = \frac{\mu}{\lambda + \mu} ; \quad p_{11}^* = \frac{\lambda}{\lambda + \mu}.$$

$$\begin{aligned} P^* &= \frac{1}{\lambda + \mu} \begin{pmatrix} \mu & \lambda \\ \mu & \lambda \end{pmatrix}. \quad (P^*)^2 = \frac{1}{(\lambda + \mu)^2} \begin{pmatrix} \mu & \lambda \\ \mu & \lambda \end{pmatrix} \begin{pmatrix} \mu & \lambda \\ \mu & \lambda \end{pmatrix} \\ &= \frac{1}{(\lambda + \mu)^2} \begin{pmatrix} \mu(\lambda + \mu) & \lambda(\lambda + \mu) \\ \mu(\lambda + \mu) & \lambda(\lambda + \mu) \end{pmatrix} = \frac{1}{\lambda + \mu} \begin{pmatrix} \mu & \lambda \\ \mu & \lambda \end{pmatrix} \\ &= P^*. \end{aligned}$$

$$\Rightarrow P^{*n} = P^* \quad \forall n \geq 1.$$

$$\begin{aligned} \Rightarrow P_{ij}(t) &= \sum_{n=0}^{\infty} \overbrace{P_{ij}^{*n}}^{=P_{ij}^* \text{ if } n \geq 1} e^{-(\lambda + \mu)t} \frac{((\lambda + \mu)t)^n}{n!} \\ &= e^{-(\lambda + \mu)t} + \sum_{n=1}^{\infty} P_{ij}^* e^{-(\lambda + \mu)t} \frac{((\lambda + \mu)t)^n}{n!} \\ &= e^{-(\lambda + \mu)t} + P_{ij}^* e^{-(\lambda + \mu)t} \underbrace{\sum_{n=1}^{\infty} \frac{((\lambda + \mu)t)^n}{n!}}_{e^{(\lambda + \mu)t} - 1} \end{aligned}$$

$$= e^{-(\lambda + \mu)t} + P_{ij}^* e^{-(\lambda + \mu)t} (e^{(\lambda + \mu)t} - 1)$$

$$\boxed{P_{ij}(t) = e^{-(\lambda + \mu)t} + P_{ij}^* (1 - e^{-(\lambda + \mu)t})}. \quad P^* = \begin{pmatrix} 0 & \frac{\lambda}{\lambda + \mu} \\ \frac{\mu}{\lambda + \mu} & 0 \end{pmatrix}$$

$$\begin{aligned}\text{Therefore, } P_{00}(t) &= e^{-(\lambda+\mu)t} + \frac{\mu}{\lambda+\mu} (1 - e^{-(\lambda+\mu)t}) \\ &= \frac{\mu}{\lambda+\mu} + \frac{\lambda}{\lambda+\mu} e^{-(\lambda+\mu)t}.\end{aligned}$$

⋮

We recover the same formulas as we found in lecture 29 from March 20, when we used the Kolmogorov backward and forward equations!

\* Please fill in the course evaluations!