Recall: Exponential distribution: 
$$X \sim Exp(x)$$
  
 $pdf$ 
 $p(x) = \begin{cases} \lambda e^{-\lambda x}, x \ge 0 \\ 0, x < 0 \end{cases}$ 
 $CDF$ 
 $F(x) = \begin{cases} 1 - e^{-\lambda x}, x \ge 0 \\ 0, x < 0 \end{cases}$ 

• Minimum of 2 independent Exp r.v.  

$$X_{1} \sim Exp(\lambda_{1}), \quad X_{2} \sim Exp(\lambda_{2}) \quad \text{independent}$$

$$X = \min(X_{1}, X_{2}), \qquad X \sim ?$$
(Rumare: in the example with the 2 tellers, X would be the l<sup>st</sup> time one of the 2 distances to leave).  
Idea: Look at the event  $E = \{X > x\}$ .  
Due can show  $X > x \iff X_{1} > x$  and  $X_{2} > x$   

$$= 7 E = \{X_{1} > x, X_{2} > x\} \qquad (x \ge 0).$$
So,  $P(X > x) = P(X_{1} > x, X_{2} > x) = P(X_{1} > x) P(X_{2} > x)$   

$$x_{1} \operatorname{indep} \notin X_{2}$$

$$= e^{-\lambda_{1}x} e^{-\lambda_{2}x} = e^{-(\lambda_{1} + \lambda_{2})x}.$$

$$= 7 F_{X}(x) = 1 - e^{-(\lambda_{1} + \lambda_{2})x} \quad \text{and} \quad X \sim Exp(\lambda_{1} + \lambda_{2})!$$

We can easily generalize. <u>Proposition</u>: let  $X_{1,...,}$   $X_n$  be a independent r.v.  $X_i^{\sim} Exp(\lambda_i)$ . Then,  $X = \min(X_{1,...,}, X_n) \sim Exp(\lambda_1 + \dots + \lambda_n)$ .

Example: Recall the previous example, with 3 customers at  
The bane, and 2 tellers. What is the expected time rutil all  
3 customers have left?  
Solution: Time with 1<sup>st</sup> leaves is usin of 2 indep Exp(
$$\lambda$$
),  
so is Exp( $2\lambda$ ) with mean  $\frac{1}{2\lambda}$ .  
. Thus, down effectively restorts, but now there are 2  
customers, so first leaves after mean time  $\frac{1}{2\lambda}$ .  
. Clove restorts, and last customer leaves after mean  
time  $\frac{1}{4}$ .  
. Probability that  $X_1 \in X_2$ , where  $X_1 \approx Exp(\lambda_1)$   
Proposition: If  $X_{1,1}X_2$  are independent  $Exp(\lambda_1)$ , thus  
 $P(X_1 \in X_2) = \frac{1}{\lambda_1 + \lambda_2}$ .  
Prof :  $P(X_1 \in X_2) = \iint_{X_1 + \lambda_2} e^{\lambda_2 x_2} dx_1 dx_2$   
 $= \int_{0}^{+\infty} \int_{\lambda_1}^{X_2} e^{-\lambda_1 x_1} e^{-\lambda_2 x_2} dx_2 dx_2$   
 $= \int_{0}^{+\infty} \lambda_2 e^{-\lambda_1 x_2} (1 - e^{-\lambda_1 x_2}) dx_2$   
 $= \int_{0}^{+\infty} \lambda_2 e^{-\lambda_2 x_2} - \int_{0}^{+\infty} \lambda_2 e^{-(\lambda_1 + \lambda_2)x_2} dx_2$ 

= 
$$L - \frac{\lambda_2}{\lambda_1 + \lambda_2} = \frac{\lambda_1}{\lambda_1 + \lambda_2}$$
  
• Sum of independent exponential r.v.  
Suppose that X, Y are independent ontimous r.v., and  
let Z = X+Y. Then,  $F_2(z) = P(Z \in z) = \int \int f_{X_1} y_{X_1} y_{X_2} y_{X_3} y_{X_$ 

$$\rightarrow \text{ let } X_{1} \sim \text{Exp}(\lambda_{1}) \text{ and } X_{2} \sim \text{Exp}(\lambda_{2}).$$
Then,  $f_{X_{1}+X_{2}}(t) = (f_{X_{1}} * f_{X_{2}}(t)) = \int_{-\infty}^{\infty} f_{X_{1}}(t-s) \cdot f_{X_{2}}(s) ds$ 

$$= \int_{0}^{t} f_{X_{1}}(t-s) \cdot f_{X_{2}}(s) ds \quad (\text{ because } f_{X_{1}}(u) = 0 \text{ for } u \in 0)$$

$$= \int_{0}^{t} \lambda_{1} e^{-\lambda_{1}(t-s)} \lambda_{2} e^{-\lambda_{2}s} ds$$
  
=  $\lambda_{1} \lambda_{2} e^{-\lambda_{1}t} \int_{0}^{t} e^{-(\lambda_{2}-\lambda_{1})s} ds$   
=  $\lambda_{1} \lambda_{2} e^{-\lambda_{1}t} \frac{(1-e^{-(\lambda_{2}-\lambda_{1})t})}{\lambda_{2}-\lambda_{1}}, \text{ with } \lambda_{2} \neq \lambda_{1}$ 

=> 
$$f_{X_1+X_2}(t) = \frac{\lambda_1}{\lambda_1-\lambda_2} \lambda_2 e^{-\lambda_2 t} + \frac{\lambda_2}{\lambda_2-\lambda_1} e^{-\lambda_1 t}$$

-Muis generalizes by induction (see Ross 5.2.4) to  

$$f_{X_{i}+\cdots+X_{n}}(t) = \sum_{i=1}^{n} C_{i,n} \lambda_{i} e^{-\lambda_{i}t}$$
  
where  $C_{i,n} = \overline{\prod} \frac{\lambda_{j}}{J_{j}-\lambda_{i}}$